



## Lipid lowering therapy, low-density lipoprotein level and risk of intracerebral hemorrhage – a meta-analysis

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1 Abstract

2 Background

3 The association of lipid lowering therapy and intracerebral haemorrhage risk is controversial.

4 Methods

5 We performed a cumulative meta-analysis of lipid lowering trials that reported intracerebral  
6 haemorrhage. Statin, fibrate, ezetimibe, PCSK9 and CETP trials were included. We explored  
7 whether the association of lipid lowering therapy and risk of intracerebral haemorrhage may  
8 vary by baseline LDL level, mean change in LDL or baseline cardiovascular risk of  
9 population.

10 Results

11 Among 39 trials (287,651 participants), lipid lowering therapy was not associated with a  
12 statistically significant increased risk of ICH in primary and secondary prevention trials  
13 combined (odds ratio, 1.12; 95% CI, 0.98 to 1.28). Lipid lowering was associated with an  
14 increased risk of ICH in secondary prevention trials (odds ratio, 1.18; 95% CI, 1.00 to 1.38),  
15 but not in primary prevention trials (odds ratio, 1.01; 95% CI, 0.78 to 1.30), but the test for  
16 interaction was not significant (P for interaction = 0.31). Meta-regression of baseline LDL or  
17 difference in LDL reduction between active and control did not explain significant  
18 heterogeneity between studies for ICH risk. Of 1,000 individuals treated for one year for  
19 secondary prevention, we estimated 9.17 (95% CI, 5.78 to 12.66) fewer ischemic strokes and  
20 0.48 (95% CI, 0.06 to 1.02) more ICH, and a net reduction of 8.69 in all stroke per 1,000  
21 person-years.

1 Conclusion

2 The benefits of lipid lowering therapy in prevention of ischemic stroke greatly exceed the risk  
3 of ICH. Concern about ICH should not discourage stroke clinicians from prescribing lipid  
4 lowering therapy for secondary prevention of ischemic stroke.

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## 1 Introduction

2 Randomised controlled trials have shown that low-density lipoprotein cholesterol (LDL-C)  
3 lowering with statins (3-hydroxy-3-methylglutaryl-coenzyme A reductase inhibitors) reduce  
4 the risk of ischaemic stroke (1), but there is controversy about whether lipid lowering agents  
5 increase the risk of intracerebral haemorrhage (ICH) (2,3). While epidemiological studies  
6 report a positive association between high serum LDL-C and ischaemic stroke (4–6), the  
7 association with LDL-C and ICH appears inverse (6–8).

8         Prior meta-analyses, evaluating the association of statin therapy and ICH, have  
9 reported no overall increase in risk of ICH (10), although one large trial reported an increased  
10 risk of ICH among those randomised to high-dose statin therapy (11). Proposed mechanisms,  
11 through which an increased risk of ICH may be mediated, include low levels of LDL-C  
12 weakening the endothelium of intracerebral arteries, causing haemorrhagic stroke in the  
13 setting of hypertension (6). Another potential mechanism is the pleiotropic  
14 antiplatelet/antithrombotic effect of lipid lowering therapies, especially statins (12). To date,  
15 meta-analyses of randomized controlled trials evaluating statin therapy, have reported on the  
16 risk of ICH, but have not explored all lipid lowering therapies and whether baseline LDL, or  
17 cardiovascular risk changes the association of all lipid lowering therapies with ICH.

18         In this meta-analysis of lipid lowering phase III trials, we sought to determine whether  
19 lipid lowering therapy increased the risk of ICH overall, and within pre-specified subgroups  
20 of participants (i.e. those with lower baseline LDL-C level, larger magnitude of LDL  
21 reduction and prior cardiovascular disease).

## 1 Methods

### 2 Cumulative meta-analysis

3 We extracted data from two previous meta-analyses: one of randomised controlled trials of  
4 statin therapy for cardiovascular prevention, reporting ICH outcomes (10) and the other of  
5 randomized controlled trials of fibrates for prevention of cardiovascular outcomes, reporting  
6 ICH (14). We limited our search to dates not included in these reviews (2012 – 2018) and  
7 repeated primary data extraction for all papers to confirm accuracy.

### 8 Selection Criteria

9 We performed a systematic review, adhering to the PRISMA guidelines (15), to select  
10 randomised controlled trials of lipid lowering therapy that reported haemorrhagic stroke on  
11 follow-up. We included all trials with: subjects > 18 years, lipid lowering therapy and  
12 haemorrhagic stroke outcome data. We limited our search to published, peer-reviewed studies  
13 in English.

### 14 Search strategy

15 We developed a search strategy for the PUBMED database. The database was searched from  
16 Jan 2012 to May 2018. Four reviewers (CJ, SR, MC, RM) independently screened titles and  
17 abstracts. Full texts were sourced for relevant articles. Inclusion criteria were assessed  
18 independently, and the final list was agreed by consensus. We also screened the reference list  
19 of similar review articles and earlier published meta-analyses obtained in our search.

### 20 Data extraction

21 For each study, we extracted the title, year of publication, active and control numbers, major  
22 bleeding and stroke outcome data. Stroke outcome was classified as either ischaemic or  
23 haemorrhagic, if available. We also collected baseline mean LDL-C, mean HDL-C and  
24 change in LDL-C from baseline to follow-up (if available). We labelled the studies as either

1 primary or secondary prevention. We used a definition of greater than fifty percent baseline  
2 cardiovascular disease (stroke, myocardial infarction) as our secondary prevention cut-off.  
3 Reviewers independently extracted data, compared for inconsistencies, and merged into a  
4 final data set.

### 5 Data synthesis and analysis

6 We present a descriptive analysis of each individual trial and summarise this analysis in both  
7 table (Supplementary Table I) and figures (Figure 1-3). We calculated odds ratio (OR) and  
8 95% confidence intervals from individual studies. Weighted pooled treatment effects were  
9 calculated using a random effects model. The variability across studies due to heterogeneity  
10 was estimated with the  $I^2$  statistic. We tested for an interaction between subgroup relative  
11 risks by dividing the difference in log relative risk by its standard error (16). Statistical  
12 analysis was performed using the Metafor package (17) on R Statistical Software (V3.4.3).

### 13 Results

14 In total, thirty-nine randomised controlled trials were eligible that recruited 287,651  
15 participants and reported 27,376 deaths, 7092 ischaemic strokes and 1035 intracerebral  
16 haemorrhages. Our updated search results found 1026 studies, 974 were excluded after title  
17 and abstract screening, 29 were excluded after full text review including 18 studies that did  
18 not report intracerebral haemorrhage, leaving 5 studies for inclusion (Supplementary Figure  
19 1). Thirty-one were trials of statins (18–28,11,29–48), four were studies of fibrates (49–52),  
20 two were studies of statins in combination with ezetimibe (50,53), one was a study of a  
21 proprotein convertase subtilisin/kexin type 9 (PCSK9) inhibitor (54) and one was a study of a  
22 cholesteryl ester transfer protein (CETP) inhibitor (55). The mean follow-up across all studies  
23 was 3.97 years. The mean age was 62.4 years in the active group and 62.4 years in the control  
24 group. 38,750 (26.9%) patients were female in the active group and 37,949 (26.4%) were  
25 female in the control group.



## 1 Meta-analysis of intracerebral haemorrhage

2 In all trials, ICH occurred in 549 (0.38%) patients in the active group and 486 (0.34%)  
3 patients in the control group (odds ratio, 1.12; 95% CI, 0.98 to 1.28) (Figure 1). The P for  
4 heterogeneity was 0.23,  $I^2=4.2\%$ ,  $Q=44.02$ , and degrees of freedom=38.

## 5 Meta-analysis of intracerebral haemorrhage (Primary and Secondary Prevention)

6 We repeated the analysis, separately, for primary and secondary prevention trials. Active  
7 therapy was associated with an increased risk for ICH in secondary prevention trials (odds  
8 ratio, 1.18; 95% CI, 1.00 to 1.38) (Figure 1). The P for heterogeneity was 0.2646,  $I^2=5.36\%$ ,  
9  $Q=24.6078$ , and degrees of freedom=21. Active therapy was not associated with an increased  
10 risk for ICH in primary prevention trials (odds ratio, 1.01; 95% CI, 0.78 to 1.30) (Figure 1).  
11 The P for heterogeneity was 0.31,  $I^2=12.44\%$ ,  $Q=18.3126$ , and degrees of freedom=16. The P  
12 for interaction was not significant (0.31).

## 13 Meta-analysis of ischaemic stroke and all-cause mortality

14 Ischaemic stroke occurred in 3,213 (2.23%) patients in the active group and 3,879 (2.7%)  
15 patients in the control group. Lipid lowering therapy was associated with a significant  
16 decrease in ischaemic stroke (odds ratio, 0.82; 95% CI, 0.76 to 0.88) and all-cause mortality  
17 (odds ratio, 0.94; 95% CI, 0.90 to 0.98). (Figure 2 and 3).

## 18 Meta-regression - baseline LDL, LDL reduction (Active) and difference in LDL reduction 19 between Active and Control

20 Three meta-regressions were performed to examine whether any between-study heterogeneity  
21 could be explained by baseline LDL-C value, mean LDL-C difference pre and post treatment  
22 and by the difference in mean LDL-C reduction between active and control treatments. The  
23 regression coefficient for baseline LDL-C was not statistically significant at 0.0005 (95% CI,  
24 -0.0044 to 0.0054,  $p=0.8323$ ,  $I^2=7.45\%$ ) (Figure 3). The regression coefficient for mean LDL-

1 C difference was small and not statistically significant at 0.0039 (95% CI, -0.0050 to  
2 0.01280,  $p=0.3914$ ,  $I^2=0.00\%$ ) (Supplementary Figure 3). The regression coefficient for  
3 difference in LDL-C reduction was small and not statistically significant at 0.0011 (95% CI, -  
4 0.0080 to 0.0103,  $p=0.8068$ ,  $I^2=7.88\%$ ) (Supplementary Figure 4).

## 5 Discussion

### 6 Main findings

7 We performed a systematic review and meta-analysis of all randomised controlled trials of  
8 lipid lowering therapy to investigate the relationship between lipid lowering and ICH. We did  
9 not find a statistically significant increased risk of ICH with lipid lowering overall (odds  
10 ratio, 1.12; 95% CI, 0.98 to 1.28), but on subgroup analysis of trials, secondary prevention  
11 was significant for lipid lowering and ICH risk in secondary prevention trials (odds ratio,  
12 1.18; 95% CI, 1.00 to 1.38), however, the P for interaction was not significant (0.31). Lipid  
13 lowering therapy was associated with a statistically significant reduced risk of ischaemic  
14 stroke (odds ratio, 0.82; 95% CI, 0.76 to 0.88). An additional meta-regression analysis was  
15 performed: baseline LDL (active), difference in LDL reduction (active) or difference in LDL  
16 reduction between active and control did not explain significant heterogeneity between  
17 studies for ICH risk.

18 Prior meta-analyses have not reported an increased risk of ICH with lipid lowering (1,10,56),  
19 but these only included statin trials. In contrast, we included trials of all lipid lowering  
20 therapies on the premise that lower LDL levels may increase the risk of ICH, as suggested by  
21 epidemiologic studies (6–8), and may not be related to a class effect. Therefore, our work  
22 builds on these studies by adding data from additional statin trials since 2012 (two), lipid  
23 lowering fibrate trials (four), PCSK9 inhibitors (one), CETP inhibitors (one) and a meta-

1 regression of baseline LDL, LDL reduction (active) and LDL reduction between active and  
2 control.

3 The ICH risk becomes more apparent with an increased event rate, this occurs in two  
4 scenarios, one, when there is a higher risk of bleeding i.e. secondary prevention higher risk  
5 population and two, large studies. Supplementary figure 5 demonstrates these two scenarios  
6 by showing a linear association between ICH event rates and ischaemic stroke rates, which is  
7 expected and consistent with other epidemiological observations and relates to common risk  
8 factors for ischemic stroke and ICH.

9 There is uncertainty and reluctance to continue lipid lowering medications immediately post  
10 acute stroke (9). The 2013 ACC/AHA guidelines only give statin prescribing a moderate IIa  
11 rating (57). To illustrate how our findings apply to everyday clinical practice, we applied the  
12 relative risk of lipid lowering on ICH (1.12) to the absolute baseline risk of ICH from the  
13 control group of our meta-analysis (0.34%). The corresponding Number Needed to Harm  
14 (NNH) for ICH with lipid lowering was 2451 (95% CI, 1158 to 20875). We then applied the  
15 relative risk of lipid lowering on ischaemic stroke (0.82) to the baseline risk of ischaemic  
16 stroke from the control group of our meta-analysis (2.7%). The corresponding Number  
17 Needed to Treat (NNT) for preventing ischaemic stroke with lipid lowering was 206 (95%  
18 CI, 150 to 328). This means, of 1,000 individuals treated for one year with lipid lowering  
19 therapy, we estimated 9.17 (95% CI, 5.78 to 12.66) fewer ischemic strokes and 0.41 (95% CI,  
20 0.05 to 0.86) more ICH, and a net reduction of 8.77 in all stroke per 1,000 person-years. We  
21 were unable to identify a clinical scenario that would discourage stroke clinicians from  
22 prescribing lipid lowering therapy.

1 **Strengths and limitations**

2 The definition of ICH varies between clinical studies and failure to classify correctly could  
3 lead to a non-differential misclassification bias. Eighteen studies did not report ICH outcome  
4 data and had to be excluded from the analysis, introducing a possible reporting bias. There  
5 was clinical heterogeneity between the participants in the selected trials, as this was not an  
6 individual participant level meta-analysis, we were unable to consider prior history of ICH.  
7 Strengths of this systematic review include the inclusion of five classes of lipid lowering  
8 drugs and subgroup analysis by prevention type and combining the relative risk reduction of  
9 lipid lowering and ICH with the absolute risk of ICH which should provide some level of  
10 reassurance to physicians with regards to the risk-benefit profile of lipid lowering in stroke  
11 patients.

12 **Implications**

13 In conclusion, lipid lowering therapy is not associated with a statistically significant  
14 increased risk of ICH overall. Baseline LDL level, change in LDL post treatment or  
15 difference in LDL reduction between active and control are not associated with a statistically  
16 significant increased risk of ICH.

17 In the general population, the benefits of lipid lowering therapy in prevention of ischemic  
18 stroke greatly exceed the risk of ICH.

1 **Contributors**

2 CJ, SR, MC, RM were responsible for data collection. CJ performed the analysis. All authors  
3 contributed to data interpretation and critical revision of the report.

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5 The corresponding author certifies that no other persons have made substantial contributions  
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14 **Disclosures**

15 All authors declare no competing interests.

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## 1 Figure Legends

### 2 Figure 1 – Lipid lowering and intracerebral haemorrhage

3 Figure 1 - Forest plot for intracerebral haemorrhage. Forest plot showing the effect of lipid lowering  
4 therapy on intracerebral haemorrhage. The forest plot is divided in two sections according to type of  
5 prevention trial a) primary b) secondary. The squares and bars represent the mean values and 95%  
6 confidence intervals of the effect sizes, while the size of the squares reflects the weight of the studies.  
7 The combined effects (sub-summary and summary) appear as diamonds and the vertical dashed line  
8 represents the line of no effect.

### 9 Figure 2 – Lipid lowering and ischaemic stroke

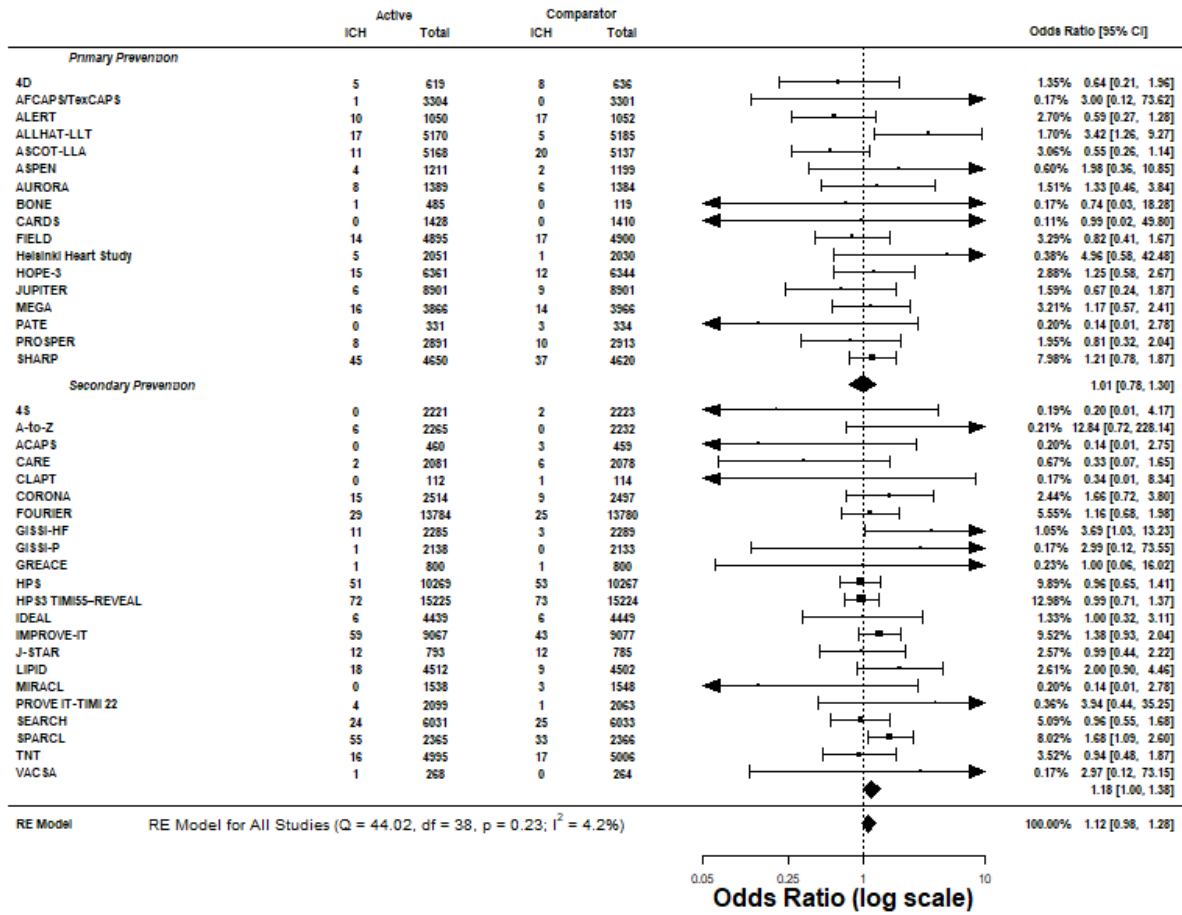
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14 effects (sub-summary and summary) appear as diamonds and the vertical dashed line represents the  
15 line of no effect.

### 16 Figure 3 – Meta-regression scatterplot – baseline mean LDL (active) and intracerebral 17 haemorrhage

18 Figure 3 – A scatterplot of the risk ratio for each study by baseline LDL cholesterol (predictor). Each  
19 study is represented by a circle. The circle sizes are proportional to the inverse of the standard errors  
20 (i.e., larger/more precise studies are shown as larger points). The solid line represents the predicted  
21 average risk ratio as a function of baseline LDL cholesterol (predictor). The dashed lines represent the  
22 95% confidence interval.

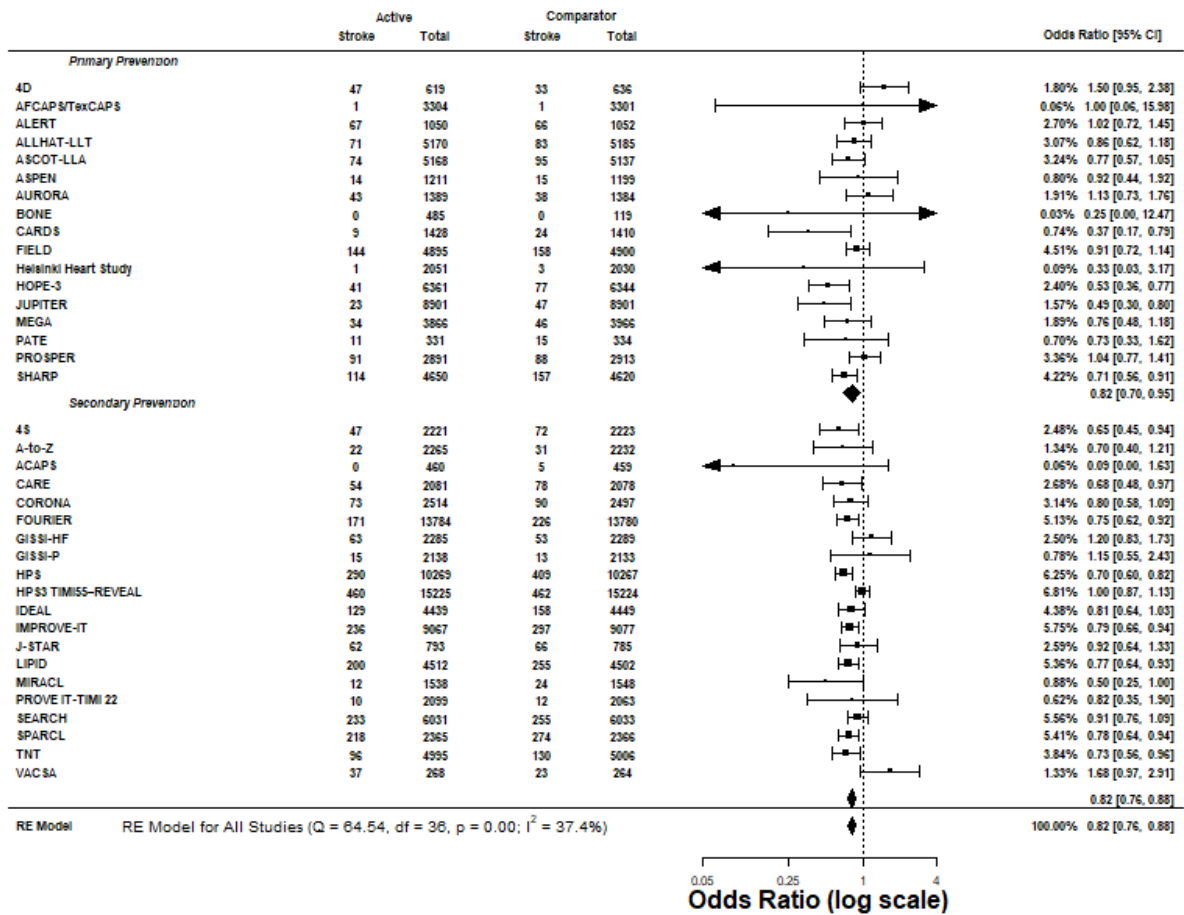
# 1 Figures

## 2 Figure 1 – Lipid lowering and intracerebral haemorrhage



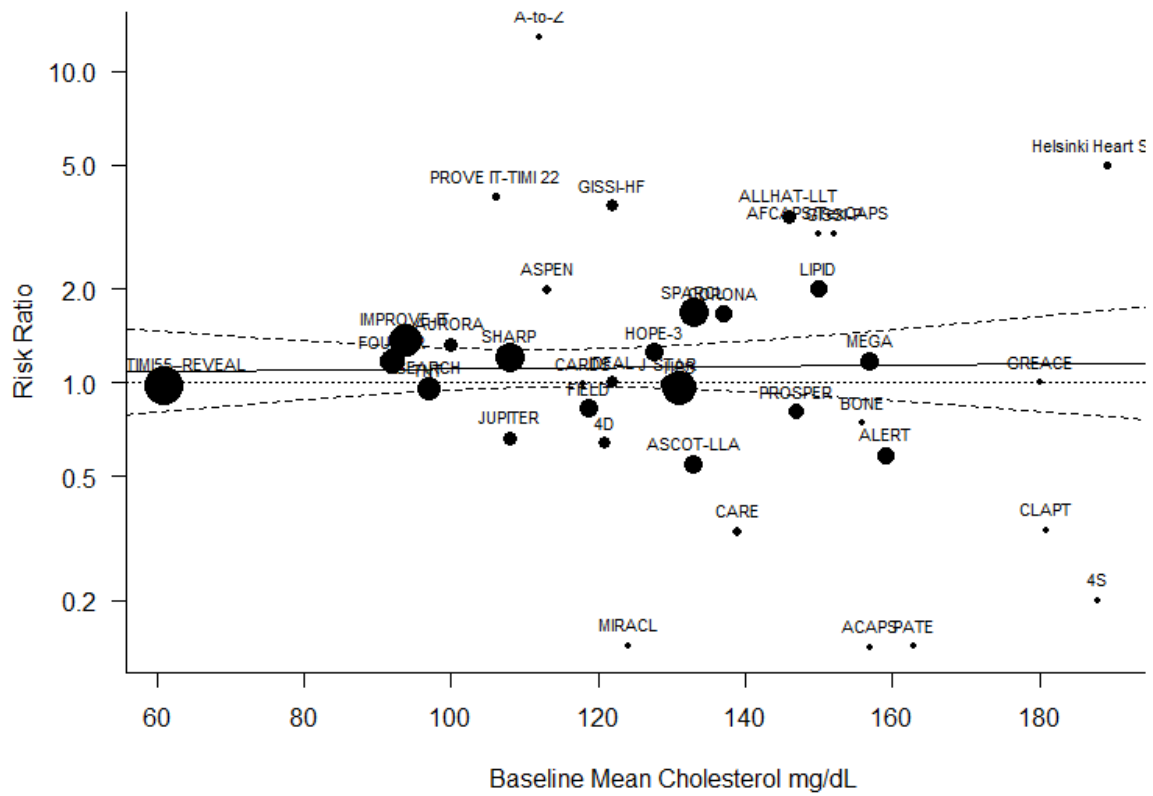
### 3

1 Figure 2 – Lipid lowering and ischaemic stroke



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- 1 Figure 3 - Meta-regression scatterplot – baseline mean LDL (active) and intracerebral
- 2 haemorrhage



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